MICROCOPY RESOLUTION TEST CHART

AD-A186 037

	READ INSTRUCTIONS BEFORE COMPLETING FORM
TION PAGE	

1. REPORT NUMBER	2. GOVT ACCESSION NO.	3. RECIPIENT'S CATALOG NUMBER

4. TITLE (and Subtitle)	5. TYPE OF REPORT & PERIOD COVERED
Necessary and sufficient conditions for the convergence of integrated and mean-integrated r-th order error of histogram density estimates	~~Technical~~ Journal – April 1987
	6. PERFORMING ORG. REPORT NUMBER 87-07

7. AUTHOR(s)	8. CONTRACT OR GRANT NUMBER(s)
X.R. Chen and L.C. Zhao	F49620-85-C-0008 ~~AIR FORCE~~

9. PERFORMING ORGANIZATION NAME AND ADDRESS	10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS
Center for Multivariate Analysis Fifth Floor Thackeray Hall University of Pittsburgh, Pittsburgh, PA 15260	61102F 2304 A5

11. CONTROLLING OFFICE NAME AND ADDRESS	12. REPORT DATE
Air Force Office of Scientific Research Department of the Air Force Bolling Air Force Base, DC 20332	April 1987
	13. NUMBER OF PAGES 16

14. MONITORING AGENCY NAME & ADDRESS(if different from Controlling Office)	15. SECURITY CLASS. (of this report)
Same as 11	Unclassified
	15a. DECLASSIFICATION/DOWNGRADING SCHEDULE

16. DISTRIBUTION STATEMENT (of this Report)

Approved for public release; distribution unlimited.

17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report)

18. SUPPLEMENTARY NOTES

19. KEY WORDS (Continue on reverse side if necessary and identify by block number)

consistency, density estimation, histogram

20. ABSTRACT (Continue on reverse side if necessary and identify by block number)

Suppose that $X_1, \ldots, X_n$ are samples drawn from a population with density function f, and $f_n(x) = f_n(x; X_1, \ldots, X_n)$ is an estimate of f(x). Denote by $m_{nr} = \int |f_n(x) - f(x)|^r dx$ and $M_{nr} = E(m_{nr})$ the Integrated r-th Order Error and Mean Integrated r-th Order Error of f_n for some $r \geq 1$ (when r = 2, they are the familiar

DD FORM 1473

20

and widely studied ISE and MISE). In this paper the same necessary and sufficient condition for $\lim_{n\to\infty} M_{nr} = 0$ and $\lim_{n\to\infty} m_{nr} = 0$ a.s. is obtained when $f_n(x)$ is the ordinary histogram estimator.

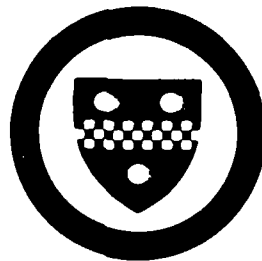| Accession For	
NTIS GRA&I	⊠
DTIC TAB	☐
Unannounced	☐
Justification	
By	
Distribution/	
Availability Codes	
Dist	Avail and/or Special
A-1	

NECESSARY AND SUFFICIENT CONDITIONS FOR
THE CONVERGENCE OF INTEGRATED AND
MEAN-INTEGRATED r-th ORDER ERROR OF
HISTOGRAM DENSITY ESTIMATES*

X.R. Chen and L.C. Zhao

Center for Multivariate Analysis
University of Pittsburgh

Center for Multivariate Analysis

University of Pittsburgh

NECESSARY AND SUFFICIENT CONDITIONS FOR
THE CONVERGENCE OF INTEGRATED AND
MEAN-INTEGRATED r-th ORDER ERROR OF
HISTOGRAM DENSITY ESTIMATES*

X.R. Chen and L.C. Zhao

Center for Multivariate Analysis
University of Pittsburgh

April 1987

Technical Report No. 87-07

Center for Multivariate Analysis
Fifth Floor Thackeray Hall
University of Pittsburgh
Pittsburgh, PA 15260

ABSTRACT

Suppose that $X_1, \ldots, X_n$ are samples drawn from a population with density function f, and $f_n(x) = f_n(x; X_1, \ldots, X_n)$ is an estimate of $f(x)$. Denote by $m_{nr} = \int |f_n(x) - f(x)|^r dx$ and $M_{nr} = E(m_{nr})$ the Integrated r-th Order Error and Mean Integrated r-th Order Error of f_n for some $r \geq 1$ (when $r = 2$, they are the familiar and widely studied ISE and MISE). In this paper the same necessary and sufficient condition for $\lim_{n \to \infty} M_{nr} = 0$ and $\lim_{n \to \infty} m_{nr} = 0$ a.s. is obtained when $f_n(x)$ is the ordinary histogram estimator.

Suppose that $X_1, \ldots, X_n$ are iid samples drawn from a d-dimensional population with density function f. Let $f_n(x) = f_n(x; X_1, \ldots, X_n)$ be an estimator of $f(x)$. The Integrated Square Error (ISE) and Mean Integrated Square Error (MISE) of f_n are defined by

$$ISE = \int |f_n(x) - f(x)|^2 dx,$$

$$MISE = E\int |f_n(x) - f(x)|^2 dx$$

respectively. They are important and widely used criteria in evaluating the performance of an estimator f_n. Quite a lot of works appeared in the statistical literature dealing with the asymptotic properties of them, for various types of estimators, such as kernel estimator, orthogonal series estimator, nearest neighbor estimator etc. For example, a much discussed problem is to find the conditions under which the assertions

$$\lim_{n\to\infty} MISE = 0, \qquad \lim_{n\to\infty} ISE = 0, a.s.$$

(1)

are true. Various conditions have been proposed in the literature which are far from necessary

In this paper we obtain the necessary and sufficient conditions of (1) for the histogram estimator In fact, we have done more For given constant $r \geq 1$, define the integrated r-th order error m_{nr} and mean integrated r-th order error M_{nr} by

$$m_{nr} = \int |f_n(x) - f(x)|^r dx, \qquad M_{nr} = E(m_{nr})$$

We obtain the necessary and sufficient conditions of $M_{nr} \to 0$ and $m_{nr} = 0$, a.s. In the case of $r=1$, the problem has been considered by Abou-Jaoude[1], [2] (see also [4], pp 19-23)

A d-dimensional histogram is defined as follows Choose $\mathbf{a}_n = (a_{1n}, \ldots, a_{dn})$ $\in R^d$, $\tilde{\mathbf{h}}_n = (h_{1n}, \ldots, h_{dn}) \in R^d$ with $h_{in} > 0$, $i=1, \ldots, d$ Denote by k the number of those $X_1, \ldots, X_n$ falling into the set $A_n(c_1, \ldots, c_d) = \{x=(x^{(1)}, \ldots, x^{(d)}) \ a_{in} + c_i h_{in} \leq x^{(i)} < a_{in} + (c_i + 1)h_{in}, i=1, \ldots, d\}$

Define

$$f_n(x) = k/(nh_{1n} \cdots h_{dn}), \quad \text{when } x \in A_n(c_1, \ldots, c_d),$$

$$c_1, \ldots, c_d = 0, \pm 1, \pm 2, \ldots$$

(2)

which is a histogram estimator of $f(x)$.

Theorem. Suppose that

$$\int f^r(x)\, dx < \infty, \quad \text{for some } r \geq 1$$

(3)

$$\lim_{n \to \infty} \max_{1 \leq i \leq d} h_{in} = 0$$

(4')

$$\lim_{n \to \infty} nh_{1n} \cdots h_{dn} = \infty$$

(5')

Then for the histogram f_n defined by (2) we have

$$\lim_{n \to \infty} m_{nr} = 0, \text{a.s.}$$

(6)

$$\lim_{n \to \infty} M_{nr} = 0,$$

(7)

Conversely, if (6) or (7) is true, then (3), (5') are true. Further, if $f(x)$ does not have a form

$$f(x) = \sum_{i_1, \cdots, i_d = -\infty}^{\infty} b_{i_1 \cdots i_d} I_{A(i_1, \ldots, i_d)}(x)$$

Then (4') is also true. Here $b_{i_1 \cdots i_d}$ is a constant, $A(i_1, \ldots, i_d) = \{x = (x^{(1)}, \ldots, x^{(d)})' : a_{j0} + i_j h_{j0} \leq x^{(j)} < a_{j0} + (i_j + 1)h_{j0}, j = 1, \ldots, d\}$. and $I_B(x)$ denotes the indicator of B.

Proof. For simplicity of writing, we shall give only the proof for the case of $d = 1$, as the general case involves no essential change. For $d = 1$ and writing h_n for $\tilde{h}_n$, the conditions (4') and (5') reduce to

$$\lim_{n \to \infty} h_n = 0$$

(4)

$$\lim_{n\to\infty} nh_n = \infty$$

<div align="right">(5)</div>

In the sequel we shall often write h for h_n, a for a_n. Note that we may assume $|a| \leq h$ without loss of generality. Write

$$\Delta_\ell = \Delta_{n\ell} = [a + \ell h, a + (\ell + 1)h),$$
$$I_\ell(x) = I_{\Delta_\ell}(x),$$
$$p_\ell = \int_{\Delta_\ell} f(x)\, dx$$

Then $Ef_n(x) = p_\ell/h$, when $x \in \Delta_\ell$. The symbol "C" will be used to denote any constant not depending on n (but may depend on r).

The proof will be divided into four parts.

(A). Sufficiency of (3)–(5) for (7)

First consider the case of $r > 1$. In order to prove (7), we need only to show that

$$\int |f_n(x) - Ef_n(x)|^r dx = n^{-r} h^{-r+1} \sum_\ell \left| \sum_{i=1}^{n} (I_\ell(X_i) - p_\ell) \right|^r \to 0$$

<div align="right">(8)</div>

$$\int |f(x) - Ef_n(x)|^r dx \to 0$$

<div align="right">(9)</div>

as $n \to \infty$. (9) is easy to prove : Suppose first that f is continuous on $(-\infty,\infty)$, then for any fixed constant $A > 0$,

$$\int_{-A}^{A} |f(x) - Ef_n(x)|^r dx \to 0, \text{ as } n \to \infty$$

On the other hand, by Hölder inequality, for any fixed constant $B > 0$ we have

$$\sum_\ell \left\{ \int_{\Delta_\ell} (Ef_n(x))^r dx : \Delta_\ell \subset [-B,B]^c \right\}$$

$$= \sum_\ell \left\{ h^{-r+1} \left(\int_{\Delta_\ell} f(x)\, dx \right)^r : \Delta_\ell \subset [-B,B]^c \right\}$$

$$\leq \sum_\ell \left\{ \int_{\Delta_\ell} f^r(x)\, dx : \Delta_\ell \subset [-B,B]^c \right\}$$

$$\leq \int_{[-B,B]^c} f^r(x)\,dx$$

<div align="right">(10)</div>

Hence

$$\sum_{\Delta_\ell} \{\int_{\Delta_\ell} |f(x) - Ef_n(x)|^r dx : \Delta_\ell \subset [-B,B]^c\}$$

$$\leq 2^r \{\int_{[-B,B]^c} f^r(x)\,dx + \sum (\int_{\Delta_\ell} (Ef_n(x))^r dx : \Delta_\ell \subset [-B,B]^c\}$$

$$\leq 2^{r+1} \int_{[-B,B]^c} f^r(x)\,dx$$

Summing up the above arguments and noticing (3), we obtain (9). For the general case, choose a function g such that $\int |f(x) - g(x)|^r dx \leq$ some given $\epsilon > 0$, define $g_n(x) = h^{-1}\int_{\Delta_\ell} g(x)dx$ for $x \in \Delta_\ell$, $\ell = 0, \pm 1, \ldots$, an argument similar to that leading to (10) gives

$$\int |Ef_n(x) - g_n(x)|^r dx \leq \int |f(x) - g(x)|^r dx \leq \epsilon$$

From this and the fact that (9) is true when f is continuous, (9) follows easily.

For a proof of (8), put $Y_{i\ell} = I_\ell(X_i) - p_\ell$, $S_{j\ell} = \sum_{i=1}^{j}(I_\ell(X_i) - p_\ell)$, $T_{nr} = \sum_\ell |S_{n\ell}|^r$. If $1 < r \leq 2$, then from the inequality $|1 + a|^r \leq 1 + ra + C|a|^r$ (a: real), we have

$$|S_{n\ell}|^r \leq |S_{n-1,\ell}|^r + C|Y_{n\ell}|^r + r|S_{n-1,\ell}|^{r-2}S_{n-1,\ell}Y_{n\ell}$$

Therefore,

$$E|S_{n\ell}|^r \leq E|S_{n-1,\ell}|^r + CE|Y_{n\ell}|^r \leq C\sum_{i=1}^{n} E|Y_{i\ell}|^r \leq 2Cnp_\ell$$

which in turn implies

$$n^{-r}h^{-r+1} \sup_{1\leq j\leq n} ET_{jr} \leq n^{-r}h^{-r+1}\sum_\ell 2Cnp_\ell = 2C(nh)^{-r+1} \to 0$$

<div align="right">(11)</div>

Suppose that (11) is true for $r \in (1, m]$, we proceed to show that it is true for $r \in (m, m+1]$. Since when $r > 2$ we have the inequality

$$|1 + a|^r \leq 1 + ra + Ca^2(1 + |a|^{r-2}) \text{ (a: real)}$$

It follows that

$$|S_{n\ell}|^r \le |S_{n-1,\ell}|^r + r|S_{n-1,\ell}|^{-2}S_{n-1,\ell}Y_{n\ell}$$
$$+ C|S_{n-1,\ell}|^{r-2}Y_{n\ell}^2 + C|Y_{n\ell}|^r$$

(12)

which in turn implies

$$E|S_{n\ell}|^r \le E|S_{n-1,\ell}|^r + CE|S_{n-1,\ell}|^{r-2}EY_{n\ell}^2 + CE|Y_{n\ell}|^r$$

Noticing that $EY_{n\ell}^2 \le p_\ell$, $E|Y_n|^r \le p_\ell$, we have

$$E|S_{\tilde n\ell}|^r \le E|S_{\tilde n-1,\ell}|^r + Cp_\ell(E|S_{\tilde n-1,\ell}|^{r-2}+1), \quad \tilde n = n, n-1, \ldots$$

Therefore

$$E|S_{n\ell}|^r \le Cp_\ell \sum_{j=1}^{n-1} E|S_{j\ell}|^{r-2} + np_\ell C$$

(13)

$$n^{-r}h^{-r+1} \sup_{1\le j\le n} ET_{jr}$$
$$\le C(nh)^{-r+1} \max_{1\le j\le n}\sum_\ell p_\ell E|S_{j\ell}|^{r-2} + C(nh)^{-r+1}$$

(14)

Since f is a probability density and $\int f^r(x)dx < \infty$, $r > 2$, we have $\int f^{r-1}(x)dx < \infty$.
Hence when $r > 2$

$$\sum_\ell h^{-r+2}p_\ell^{r-1} \le \int f^{r-1}(x)dx = C < \infty$$

(15)

By induction hypothesis

$$n^{-r+1}h^{-r+2} \max_{1\le j\le n} ET_{j,r-1} \to 0, \quad \text{as } n \to \infty$$

(16)

from (14)-(16), we have

$$n^{-r}h^{-r+1} \max_{1\le j\le n} ET_{jr}$$
$$\le C(nh)^{-r+1}(\sum_\ell p_\ell^{r-1})^{1/(r-1)} (\max_{1\le j\le n}\sum_\ell E|S_{j\ell}|^{r-1})^{(r-2)/(r-1)}$$
$$+ C(nh)^{-r+1}$$

$$\le C(nh)^{-r+1}(h^{(r-2)/(r-1)})(n^{r-2}h^{(r-2)^2/(r-1)})$$
$$(n^{-r+1}h^{-r+2}\max_{1\le j\le n}ET_{j,r-1})^{(r-2)/(r-1)} + C(nh)^{-r+1}$$

$$= C(nh)^{-1}(n^{-r+1}h^{-r+2}\max_{1\le j\le n}ET_{j,r-1})^{(r-2)/(r-1)} + C(nh)^{-r+1} \to 0$$

as $n \to \infty$. This shows that (11) is true for $r \in [m,m+1]$, concluding the proof of (11)

hence (7), for $r > 1$. When $r = 1$, (7) is a consequence of (6) for $r = 1$, which we are now going to prove.

(B) Sufficiency of (3)–(5) for (6)

Again first consider the case of $r > 1$. Define T_{nr} as before, then

$$I_{n\ell} \equiv \int |f_n(x) - f(x)|^r dx = n^{-r} h^{-r+1} T_{nr}$$

Hence we need only to show that

$$\lim_{n \to \infty} I_{nr} = 0, \quad \text{a. s.}$$

(17)

Define

$$\xi_{j\ell 1} = Y_{j\ell}, \qquad \eta_{j\ell 1} = |S_{j\ell}|^{r-1} \text{sign}(S_{j\ell})$$

$$\xi_{j\ell 2} = Y_{j\ell}^2 - p_\ell(1-p_\ell), \qquad \eta_{j\ell 2} = |S_{j\ell}|^{r-1}$$

and proceed to show that for any given $\varepsilon > 0$

$$J \equiv P(n^{-r} h^{-r+1} \max_{2 \le k \le n} |\sum_{j=2}^{k} \sum_\ell \xi_{j\ell i} \eta_{j-1,\ell i}| \ge \varepsilon) \le C n^{-2}, \quad i = 1, 2$$

(18)

In the following we shall write $\xi_{j\ell}, \eta_{j\ell}$ for $\xi_{j\ell i}, \eta_{j\ell i}$. Since $\{\sum_{j=2}^{k} \xi_{j\ell} \eta_{j-1,\ell}, k = 2, 3, \ldots\}$ is a martingale sequence, we have

$$J \le \varepsilon^{-4} n^{-4r} h^{-4r+4} CE(\sum_{j=2}^{n} \sum_\ell \xi_{j\ell} \eta_{j-1,\ell})^4$$

(19)

From an inequality of Rosenthal (see [5], p 23),

$$E(\sum_{j=2}^{n} \sum_\ell \xi_{j\ell} \eta_{j-1,\ell})^4$$

$$\le C(E\{\sum_{j=2}^{n} E[(\sum_\ell \xi_{j\ell} \eta_{j-1,\ell})^2 \mid F_{j-1}]\}^2$$

$$+ \sum_{j=2}^{n} E(\sum_\ell \xi_{j\ell} \eta_{j-1,\ell})^4)$$

(20)

Here F_j is the σ-field generated by $X_1, \ldots, X_j$. Since $\sum_\ell |\xi_{j\ell}| \le 3$, by Jensen's inequality we have

$$\sum_{j=2}^{n} E \left(\sum_{\ell} \xi_{j\ell} \eta_{j-1,\ell} \right)^4$$

$$\leq C \sum_{j=2}^{n} E \sum_{\ell} |\xi_{j\ell}| \eta_{j-1,\ell}^4$$

$$\leq C \sum_{j=2}^{n} p_\ell E \eta_{j-1,\ell}^4$$

It can be shown easily by using (13) that

$$E \eta_{j-1,\ell}^4 \leq C (1+(np_\ell)^{2r-2}) \leq C (1+(np_\ell)^{2r-1})$$

Noticing that $\sum_\ell p_\ell h^{-r+1} \leq \int |f(x)| \, dx < \infty$ and $nh \to \infty$, we have

$$\sum_{j=2}^{n} E \left(\sum_{\ell} \xi_{j\ell} \eta_{j-1,\ell} \right)^4$$

$$\leq Cn \left(\sum_{\ell} p_\ell + n^{2r-1} \sum_{\ell} p_\ell^{2r} \right)$$

$$\leq Cn \left(1 + n^{2r-1} h^{2r-2} \left(\sum_{\ell} p_\ell^r h^{-r+1} \right)^2 \right)$$

$$\leq Cn^2 (nh)^{2r-2}$$

$$(21)$$

$$E \left(\sum_{j=2}^{n} E \left[\left(\sum_{\ell} \xi_{j\ell} \eta_{j-1,\ell} \right)^2 \Big| F_{j-1} \right] \right)^2$$

$$\leq n \sum_{j=2}^{n} E \left(E \left[\left(\sum_{\ell} \xi_{j\ell} \eta_{j-1,\ell} \right)^2 \Big| F_{j-1} \right] \right)^2$$

$$\leq 2n \sum_{j=2}^{n} E \left(\sum_{\ell} p_\ell \eta_{j-1,\ell}^2 \right)^2$$

$$+ 2n \sum_{j=2}^{n} E \left(\sum_{\ell} \sum_{m} E \xi_{j\ell} \xi_{jm} \eta_{j-1,\ell} \eta_{j-1,m} \Big| F_{j-1} \right)^2$$

$$\leq Cn \sum_{j=2}^{n} E \left(\sum_{\ell} p_\ell \eta_{j-1,\ell}^2 \right)^2 + Cn \sum_{j=2}^{n} E \left(\sum_{\ell} p_\ell |\eta_{j-1,\ell}| \right)^4$$

$$\leq Cn \sum_{j=2}^{n} E \left(\sum_{\ell} p_\ell \eta_{j-1,\ell}^2 \right)^2$$

$$(22)$$

The last step is valid in view of Jensen's inequality and $\sum_\ell p_\ell = 1$. Therefore, from

$$E \eta_{j-1,\ell}^2 \eta_{j-1,m}^2 \leq (E \eta_{j-1,\ell}^4 E \eta_{j-1,m}^4)^{1/2},$$

and $\quad E \eta_{j-1,\ell}^4 \leq C(1+(np_\ell)^{2r-2})$

We obtain

$$E\left(\sum_{j=2}^{n} E\left[\left(\sum_{\ell} \xi_{j\ell}\eta_{j-1,\ell}\right)^2 \Big| F_{j-1}\right]\right)^2$$

$$\leq Cn\sum_{j=2}^{n}\sum_{\ell} p_{\ell}^2 E\eta_{j-1,\ell}^4 + Cn\sum_{j=2}^{n}\sum_{\ell\neq m} p_{\ell}p_{m}(E\eta_{j-1,\ell}^4 E\eta_{j-1,m}^4)^{1/2}$$

$$\leq Cn\sum_{j=2}^{n}\sum_{\ell} p_{\ell}^2(1+(np_{\ell})^{2r-2}) + Cn\sum_{j=2}^{n}\left(\sum_{\ell} p_{\ell}(1+(np_{\ell})^{r-1})\right)^2$$

$$\leq Cn^2(1+(nh)^{2r-2}(\sum_{\ell} p_{\ell}^r h^{-r+1})^2) + Cn^2(1+(nh)^{r-1}\sum_{\ell} p_{\ell}^r h^{-r+1})^2$$

$$\leq Cn^2(nh)^{2r-2}$$

$$\tag{23}$$

From (19)–(21) and (23), (18) follows.

Now to prove that for $r > 1$ and given $\epsilon > 0$, we have

$$P(n^{-r}h^{-r+1}\max_{1\leq k\leq n}T_{kr} \geq \epsilon) \leq C/n^2$$

$$\tag{24}$$

First suppose that $1 < r \leq 2$. We have

$$|S_{n\ell}|^r \leq |S_{n-1,\ell}|^r + C|Y_{n\ell}|^r + r|S_{n-1,\ell}|^{r-1}\mathrm{sign}(S_{n-1,\ell})Y_{n\ell}$$

Hence

$$|S_{n\ell}|^r \leq C\sum_{j=1}^{n}|Y_{j\ell}|^r + r\sum_{j=2}^{n}Y_{j\ell}\eta_{j-1,\ell 1}$$

and

$$n^{-r}h^{-r+1}\max_{1\leq k\leq n}T_{kr}$$

$$\leq Cn^{-r}h^{-r+1}\sum_{j=1}^{n}\sum_{\ell}|Y_{j\ell}|^r$$

$$+ n^{-r}h^{-r+1}r_2\max_{\leq k\leq n}\left|\sum_{j=2}^{k}\sum_{\ell}Y_{j\ell}\eta_{j-1,\ell 1}\right|$$

$$\tag{25}$$

Since

$$n^{-r}h^{-r+1}\sum_{j=1}^{n}\sum_{\ell}|Y_{j\ell}|^r \leq 2(nh)^{-r+1} \to 0$$

From (18) and (25), (24) follows.

Suppose that (24) is true for $r \leq m$, and proceed to show that it is true for $r \in (m, m+1]$. Since $r > 2$, we have

$$|S_{n\ell}|^r \le |S_{n-1,\ell}|^r + r|S_{n-1,\ell}|^{r-2}S_{n-1,\ell}Y_{n\ell}$$
$$+ C|S_{n-1,\ell}|^{r-2}Y_{n\ell}^2 + C|Y_{n\ell}|^r$$

Hence

$$n^{-r}h^{-r+1} \max_{1\le k\le n} T_{kr} \le Cn^{-r}h^{-r+1} \sum_{i=1}^{n} \sum_{\ell} |Y_{j\ell}|^r$$

$$+ rn^{-r}h^{-r+1} \max_{2\le k\le n} |\sum_{j=2}^{k} \sum_{\ell} Y_{j\ell}\eta_{j-1,\ell}|$$

$$+ Cn^{-r}h^{-r+1} \max_{2\le k\le n} |\sum_{j=2}^{k} \sum_{\ell} \xi_{j\ell 2}\tilde{\eta}_{j-1,\ell}|$$

$$+ C(nh)^{-r+1} \max_{2\le j\le n} \sum_{\ell} p_\ell |S_{j-1,\ell}|^{r-2} \equiv \sum_{i=1}^{4} H_i$$

$$(26)$$

where $\tilde{\eta}_{j-1,\ell} = |S_{j-1,\ell}|^{r-2}$. Observe that H_2/r is the expression in (18), nhH_3/C is also of the form in (18) with r replaced by $r-1$. Further, since $r > 2$ and $\int f^r(x)dx < \infty$, we have $\int f^{r-1}(x)dx < \infty$. Therefore, the inequality (18) can be applied to both H_2 and nhH_3. Further, $nh \to \infty$, so we obtain

$$P(H_2 + H_3 \ge \epsilon) \le C/n^2$$

$$(27)$$

for arbitrarily given $\epsilon > 0$. Also,

$$n^{-r}h^{-r+1} \sum_{j=1}^{n} \sum_{\ell} |Y_{j\ell}|^r \le 2(nh)^{-r+1} \to 0$$

$$(28)$$

Since $\int f^{r-1}(x)dx < \infty$, by induction hypothesis, we have

$$P(n^{-r+1}h^{-r+2} \max_{1\le k\le n} T_{k,r-1} \ge \epsilon) \le C/n^2$$

$$(29)$$

By Hölder inequality and the fact $\sum_{\ell} p_\ell^{r-1} \le Ch^{r-2}$,

$$(nh)^{-r+1} \max_{2\le j\le n} \sum_{\ell} p_\ell |S_{j-1,1}|^{r-2}$$
$$\le C(nh)^{-1} (n^{-r+1}h^{-r+2} \max_{1\le k\le n} T_{k,r-1})^{(r-2)/(r-1)}$$

$$(30)$$

(29), (30) together give

$$P(H_4 \geq \epsilon) \leq C/n^2$$

(31)

for arbitrarily given $\epsilon > 0$. Summing up (26)–(28) and (31) we see that (24) is true for $r \in (m, m + 1]$, concluding the proof of (24) hence (17). Thus we have proved (6) for the case of $r > 1$.

Now consider the case of $r = 1$. Since $f_n(x)$ as a function of x in $(-\infty, \infty)$ is a probability density, in order to prove (6), it suffices to show that for any fixed positive integer N, it is true that

$$\lim_{n \to \infty} \int_I |f_n(x) - f(x)| \, dx = 0, \quad a. s.$$

(32)

where $I = \bigcup_{\ell = -NH}^{NH} \Delta_\ell$, $H = [h^{-1}]$, the integer part of h^{-1}.

To prove (32), it suffices to verify the following two assertions:

$$\lim_{n \to \infty} \int_I |f(x) - Ef_n(x)| \, dx = 0$$

(33)

$$\lim_{n \to \infty} \int_I |f_n(x) - Ef_n(x)| \, dx = 0, \quad a. s.$$

(34)

The second assertion follows directly from Lemma 3 of Devroye [3] if we note the fact that $NH/n \leq N/nh \to 0$. The first assertion can be verified by using a continuous function g on $[a-N-1, a+N+1]$ such that the integral

$$\int_{a-N-1}^{a+N+1} |f(x) - g(x)| \, dx$$

does not exceed given $\epsilon > 0$. Trivial details are omitted.

(C). Necessity of (3), (4)

Since

$$\int f_n^r(x) \, dx = n^{-r} h^{-r+1} \sum_\ell |\sum_{i=1}^n (I_\ell(X_i) - p_\ell)|^r$$

$$\leq n^{-r} h^{-r+1} (\sum_\ell |\sum_{i=1}^n (I_\ell(X_i) - p_\ell)|)^r \leq 2^r/h^{r-1} < \infty$$

Therefore, if (6) or (7) is true, then (3) is true.

Suppose that $f(x)$ does not have a form

$$\sum_{\ell=-\infty}^{\infty} C_{\ell} \, I_{[a_0+\ell h_0, \, a_0+(\ell+1)h_0)}(x)$$

(35)

for some a_0, $h_0 > 0$ and $\{C_{\ell}\}$. We want to prove that if (6) or (7) is true, then $h_n \to 0$. Suppose in the contraty that $h_n \searrow 0$. Then there exists subsequence $\{h_{n_i}, i = 1, 2, \ldots\}$ such that $h_{n_i} \to h_0 > 0$ as $i \to \infty$. h_0 must be finite, otherwise we would have $f_n(x) \to 0$ uniformly in $\{X_1, X_2, \ldots\}$ and x, and this contradicts obviously with any one of (6) or (7). Since $|a_{n_i}| \leq h_{n_i}$, without losing generality we may assume that $a_{n_i} \to a_0$, also finite. From these facts it follows by the law of large numbers that if we define

$$g(x) = h_0^{-1} \int_{a_0+\ell h_0}^{a_0+(\ell+1)h_0} f(t)\,dt,$$

$$x \in [a_0+\ell h_0, \, a_0+(\ell+1)h_0), \quad \ell = 0, \pm 1, \pm 2, \ldots$$

Then we have

$$\lim_{n \to \infty} \int_I |f_n(x) - g(x)|^r dx = 0, \quad \text{a. s.}$$

(36)

for any bounded interval I. Since at least one of (6) and (7) is true, (36) implies that $f = g$ almost everywhere on $(-\infty, \infty)$ (Lebesgue measure). Hence f can be expressed in the form (35), contradicting the assumption.

(D). Necessity of (5).

Suppose in the contratry that (5) is not true, then there exists subsequence $\{h_n\}$ such that $h_{n_i} \to 0$, $n_i h_{n_i} \to u < \infty$. We restrict ourselves to the discussion of the subsequence. For simplicity of writing and without losing generality, we may assume that $h_n \to 0$ and

$$\lim_{n \to \infty} n h_n = u < \infty$$

(37)

Since $h \to 0$, we have

$$\lim_{n \to \infty} \int |f(x) - Ef_n(x)|\,dx = 0$$

(38)

In fact, as pointed out earlier, (33) is true for any bounded interval I. Since $f(x)$ and $Ef_n(x)$ are probability density functions on $(-\infty, \infty)$, this fact implies (38). Now it is easily seen that if at least one of (6) and (7) is true, then

$$\lim_{n\to\infty} E\int |f_n(x) - Ef_n(x)|dx = 0$$

(39)

In fact, if (7) is true, then for any bounded interval I we have by Hölder inequality that

$$E\int_I |f_n(x) - f(x)|dx \to 0, \quad \text{as } n \to \infty$$

(40)

Since $f(x)$, $f_n(x)$ are probability density functions, it is easily seen that

$$\int |f_n(x) - f(x)|dx \le 2\int_I |f_n(x) - f(x)|dx + 2[1-\int_I f(x)dx]$$

(41)

From (40), (41), it follows that

$$\lim_{n\to\infty} E\int |f_n(x) - f(x)|dx = 0$$

(42)

Now (39) follows from (38) and (42) If (6) is true, then by Hölder inequality we have

$$\lim_{n\to\infty} \int_I |f(x) - f_n(x)|dx = 0, \quad \text{a. s.}$$

(43)

for any bounded interval I. From (43) and the dominated convergence theorem, we again have (40) and hence (39).

First we assume that $u=0$. Write $A_n = \bigcup_{i=1}^{n} [X_i -h, X_i +h]$. By the definition, $f_n(x) = 0$ for $x \bar{\varepsilon} A_n$. So we have

$$E\int_{A_n} f_n(x)dx = 1.$$

Denote by $\lambda(A_n)$ the Lebesgue measure of A_n. Then

$$\lambda(A_n) \le 2nh \to 0, \quad \text{as } n \to \infty,$$

which implies that

$$\lim_{n\to\infty} E\int_{A_n} f(x)dx = 0.$$

Thus

$$E\int |f_n(x) - f(x)|dx \ge E\int_{A_n} |f_n(x) - f(x)|dx$$

$$\geq E \int_{A_n} f_n(x)\,dx - E \int_{A_n} f(x)\,dx \to 1, \quad \text{as } n \to \infty.$$

From this and (38), it follows that

$$\lim_{n\to\infty} E\int |f_n(x) - Ef_n(x)|\,dx = 1,$$

$$(44)$$

which contradicts (39).

Now we assume $0 < u < \infty$. By (39), there exist a sequence of positive constants, say C_n, such that $C_n \to \infty$, $C_n/n \to 0$ and

$$C_n E\int |f_n(x) - Ef_n(x)|\,dx \to 0$$

$$(45)$$

Write $k = [n/C_n]$. Then

$$C_n E\int |f_n(x) - Ef_n(x)|\,dx$$

$$= C_n(nh)^{-1} \sum_{\ell} E\left| \sum_{i=1}^{n} (1_\ell(X_i) - p_\ell) \right|$$

$$\geq (kh)^{-1} \sum_{\ell} E\left| E\left\{ \sum_{i=1}^{n} (1_\ell(X_i) - p_\ell) \mid X_1, \ldots, X_k \right\} \right|$$

$$= (kh)^{-1} \sum_{\ell} E\left| \sum_{i=1}^{k} (1_\ell(X_i) - p_\ell) \right|.$$

$$(46)$$

Since $h \to 0$ and $kh \leq nh/C_n \to 0$, by (44) we have

$$\lim_{k\to\infty} (kh)^{-1} \sum_{\ell} E\left| \sum_{i=1}^{k} (1_\ell(X_i) - p_1) \right| = 1.$$

On the other hand, (45) and (46) together imply

$$(kh)^{-1} \sum_{\ell} E\left| \sum_{i=1}^{k} (1_\ell(X_i) - p_\ell) \right| \to 0, \quad \text{as } n \to \infty.$$

Thus a contradiction is reached, and the proof of the Theorem is completed

Remark. The method of proof in (D) can easily be adopted to the case of the kernel estimate.

REFERENCES

1. Abou-Jaoude, S. (1976). Sur une condition nécessaire et suffisante de L_1-convergence presque complète de l'estimateur de la partition fixe pour une densité. Comptes Rendus de l'Académie des Sciences de Paris Série A, **283**, pp. 1107-1110.

2. Abou-Jaoude, S. (1976). Conditions nécessaires et suffisantes de convergence L_1 en probabilité de l'histogramme pour une densité. Annales de l'Institut Henri Poincaré, **12**, pp. 213-231

3. Devroye, L. (1983). The equivalence of weak, strong and complete convergence in L_1 for kernel density estimates. Ann. Statist., **11**, pp. 896-904

4. Devroye, L. and Györfi, L. (1985) Nonparametric Density Estimation The L_1 View. Wiley, New York.

5. Hall, P. and Heyde, C.C. (1980) Martingale Limit Theory and Its Application. Academic Press, New York

END

DATE
FILMED

DEC.
1987